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Fixed Points of the Smoothing Transformation

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Summary. Let $W_1, ..., W_N$ be N nonnegative random variables and let \mathfrak{M} be the class of all probability measures on $[0, \infty)$. Define a transformation Ton \mathfrak{M} by letting $T\mu$ be the distribution of $W_1X_1 + ... + W_NX_N$, where the X_i are independent random variables with distribution μ , which are independent of $W_1, ..., W_N$ as well. In earlier work, first Kahane and Peyriere, and then Holley and Liggett, obtained necessary and sufficient conditions for T to have a nontrivial fixed point of finite mean in the special cases that the W_i are independent and identically distributed, or are fixed multiples of one random variable. In this paper we study the transformation in general. Assuming only that for some $\gamma > 1$, $EW_i^{\gamma} < \infty$ for all i, we determine exactly when T has a nontrivial fixed point (of finite or infinite mean). When it does, we find all fixed points and prove a convergence result. In particular, it turns out that in the previously considered cases, T always has a nontrivial fixed point. Our results were motivated by a number of open problems in infinite particle systems. The basic question is: in those cases in which an infinite particle system has no invariant measures of finite mean, does it have invariant measures of infinite mean? Our results suggest possible answers to this question for the generalized potlatch and smoothing processes studied by Holley and Liggett.

1. Introduction

A number of authors have studied interacting particle systems in which the state of each site $x \in \mathbb{Z}^d$ is described by a nonnegative real number ([5, 11, 14], for example), or more generally the state of the system is a random measure on

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 R^d ([1, 2, 7, 8, 10, 15]). In most cases it has only been possible to find sufficient conditions for the existence and for the nonexistence of invariant measures of finite mean. Once there is one invariant measure of finite mean there is normally at least a one parameter family of them indexed by the mean. A familiar example is "independent motions" where the invariant measures are Poisson processes with mean measures λdx , $0 \le \lambda < \infty$.

As in the last example, the existence of a one parameter family of invariant measures with finite mean usually makes it possible to rule out the existence of nontrivial invariant measures with infinite mean. The idea is to show that any extremal invariant with infinite mean must be stochastically larger than all the members of the known one parameter family and hence must be the pointmass on the configuration which is identically infinity. When there is no invariant measure of finite mean (e.g. branching random walks in $d \le 2$), it is natural to ask whether the process has invariant measures of infinite mean. However, the ideas used in the finite mean case simply do not shed much light on this. In most cases, it has been difficult even to formulate a reasonable conjecture concerning this question. Examples of situations in which this question has been raised either explicitly or implicitly are: (i) open problem b of Sect. 8 of Holley and Liggett (1981) for generalized potlatch and smoothing processes, (ii) conjecture 1.6 of Liggett (1978) for independent particle systems, (iii) Theorem 1.2 of Kallenberg (1977) for critical cluster fields, and (iv) Theorem 3.1 of Dawson (1977) for measure diffusion processes. (In the latter two cases, the question is implicit in the finite intensity assumptions in the theorems. In Dawson's theorem, this assumption appears in the proof rather than in the statement of the theorem.)

This paper is an attempt to shed some light on the above question by considering a simplified version of the generalized smoothing process of Holley and Liggett (1981), which we will call the smoothing transformation. In order to define the smoothing transformation fix N nonnegative random variables $W_1, ..., W_N$ with $P(W_i > 0) > 0$, but which otherwise have an arbitrary joint distribution, and let \mathfrak{M} be the class of all probability measures on $[0, \infty)$. The smoothing transformation T on \mathfrak{M} is defined by letting $T\mu$ be the distribution of $W_1X_1 + ... + W_NX_N$, where $X_1, ..., X_N$ are independent random variables with distribution μ , and are independent of $(W_1, ..., W_N)$. Of course, T can be regarded as a (nonlinear) transformation on the class \mathfrak{L} of Laplace transforms φ of elements of \mathfrak{M} . With this interpretation, T takes the form

$$(T\varphi)(\theta) = E \prod_{i=1}^{N} \varphi(\theta W_i).$$

Kahane and Peyriere (1976) gave necessary and sufficient conditions for T to have a fixed point of finite mean in the case that W_1, \ldots, W_N are independent and identically distributed. Their work was motivated by questions raised by Mandelbrot relating to a model for turbulence. Holley and Liggett (1981) did the same in the case that W_1, \ldots, W_N are constant multiples of a fixed random variable. In their case, the problem was motivated by the connection between the smoothing transformation T and the generalized smoothing process.

In this paper we will study the transformation in general assuming only that for some $\gamma > 1$, $EW_i^{\gamma} < \infty$ for all *i* and will determine when *T* has fixed points of finite or infinite mean. To state our results we will need a number of definitions. Let \mathfrak{F} be the set of all nontrivial fixed points of T:

$$\mathfrak{F} = \{ \mu \in \mathfrak{M} : T\mu = \mu \text{ and } \mu \neq \delta_0 \}.$$

The condition for the existence of a nontrivial fixed point for T is given in terms of the function

$$v(\alpha) = \log \left(\sum_{i=1}^{N} E(W_i^{\alpha}; W_i > 0) \right)$$

which is defined for all $\alpha \ge 0$ and is finite for $0 \le \alpha \le \gamma$.

Theorem 1. $\mathfrak{F} \neq \emptyset$ if and only if for some $\alpha \in (0, 1]$, $v(\alpha) = 0$ and $v'(\alpha) \leq 0$.

Although the statement of this result is a simple dichotomy, its proof is not a two part affair – it is a combination of five propositions: (2.7), (2.12), (3.1), (3.2), and (3.5). Perhaps the best way of indicating the elements of the proof and how they fit together is to outline the argument for the example which led us to the theorem stated above.

Suppose N=2 and $W_1=W_2$ with distribution given by

$$P(W_i = A) = p$$
, $P(W_i = A^{-1}) = 1 - p$

for some A > 1 and $p \in (0, 1)$. In this example

$$v(\alpha) = \log(2[pA^{\alpha} + (1-p)A^{-\alpha}])$$

so by Theorem 1, $\Re \neq \emptyset$ if and only if for some $\alpha \in (0, 1]$,

(a)
$$u(\alpha) \equiv 2\lceil p A^{\alpha} + (1-p)A^{-\alpha} \rceil = 1$$

and

(b)
$$v'(\alpha) = \frac{u'(\alpha)}{u(\alpha)} = \frac{2(pA^{\alpha} - (1-p)A^{-\alpha})\ln A}{2(pA^{\alpha} + (1-p)A^{-\alpha})} \le 0.$$

This conclusion was arrived at by considering four cases.

Case I. v(1) = 0. In this case

$$E(W_1 + W_2) = 2[pA + (1-p)A^{-1}] = u(1) = 1$$

so a result of Holley and Liggett (1978) (see Theorem 7.1 on p. 190) can be applied to conclude that there are fixed points of finite mean if and only if

$$E(2W_i \log (2W_i)) < -2(\frac{1}{2}\log \frac{1}{2}) = \log 2$$

(the analogous transformation in their setting is $TX = W(X_1 + X_2)/2$). If we observe that $E(2W_i) = 1$ we can write the condition above as $2E(W_i \log W_i) < 0$ or in view of Theorem 1 and the fact that

$$v'(1) = u'(1) = 2(pA \ln A + (1-p)A^{-1} \ln A^{-1})$$

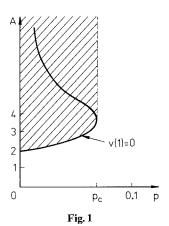
we can write the condition as v'(1) < 0.

In the last simplification we have relied on knowing the solution of the problem. Even before solving the problem however, it is easy to see that the transition between existence and nonexistence of fixed points of finite mean occurs at a special point.

In order for the equality v(1)=0 above to hold, we must have $2pA^2-A+2(1-p)=0$, i.e.

$$A = \frac{1 \pm \sqrt{1 - 16p(1 - p)}}{4p}.$$

In order for this to be real and larger than one, we must have $p \le p_c = [$ the smaller root of $16p(1-p)] = (2-\sqrt{3})/4 \approx 0.067$. When $p < p_c$ there are two possible values of A. Checking the value of $v'(\alpha)$ one finds that the smaller value of A always has $v'(\alpha) < 0$ and the larger value always has $v'(\alpha) > 0$ so the cutoff corresponds to the point on the curve $2(pA + (1-p)A^{-1}) = 1$ which is above p_c . This curve is sketched in Fig. 1.



Case II. (3.1) Suppose $v(\alpha) = 0$ and $v'(\alpha) < 0$ for some $\alpha \in (0, 1)$. Then $\mathfrak{F} \neq \emptyset$ (this covers the shaded region in Fig. 1). Define $\tilde{W_i} = W_i^{\alpha}$ and let \tilde{v} , \tilde{T} , and $\tilde{\mathfrak{F}}$ be what results when we replace W_i by $\tilde{W_i}$. It is easy to check that $\tilde{v}(1) = 0$ and $\tilde{v}'(1) < 0$ so by Case I, there is a $\psi \in \tilde{\mathfrak{F}}$ with finite mean. Let $\varphi(\theta) = \psi(\theta^{\alpha})$. If we let X(t) be the one sided stable process with index α and let τ be an independent random variable with Laplace transform ψ then $\varphi(\theta) = E \exp(-\theta X(\tau))$ so φ is a Laplace transform. It is easy to check that $T\varphi = \varphi$.

Case III. (3.5) Suppose that $v(\alpha)=0$ and $v'(\alpha)=0$ for some $\alpha \in (0,1]$. Then $\mathfrak{F} \neq \emptyset$ (this covers the dotted line at the border of the shaded region in Fig. 1). Since v is strictly convex, $v(\beta)>0$ and $v'(\beta)<0$ for all $\beta<\alpha$. Let $W_{i,\beta}=W_i\exp(-v(\beta)/\beta)$ and v_β , T_β and T_β be what results when we replace W_i by $W_{i,\beta}$. Then $v_\beta(\beta)=0$

and $v_{\beta}'(\beta) < 0$ so T_{β} has a nontrivial fixed point ψ_{β} by Case II. It turns out (see (3.2)) that $\psi_{\beta}(\infty) = t_0$ is independent of β so by scaling we can choose the fixed point ψ_{β} so that $\psi_{\beta}(1) = (t_0 + 1)/2$ and hence we can let $\beta \uparrow \alpha$ to construct a nontrivial fixed point for T.

Case IV. (2.12a) If $\mathfrak{F} \neq \emptyset$ then there is an $\alpha \in (0, 1]$ so that $v(\alpha) = 0$. (This shows that Cases I-III give us all the cases in which fixed points exist.) Let $D_{\alpha}(x) = e^{\alpha x}(1 - \varphi(e^{-x}))$. If we let Y_{α} be a random variable with distribution given by

$$Ef(Y_{\alpha}) = e^{-v(\alpha)} \sum_{i=1}^{N} E[f(-\log W_i) W_i^{\alpha}; W_i > 0]$$

and let

$$G_{\alpha}(x) = e^{\alpha x} E\left(\prod_{i=1}^{N} \varphi(e^{-x} W_i) - 1 + \sum_{i=1}^{N} (1 - \varphi(e^{-x} W_i))\right),$$

then a little calculation (see (2.3)) shows that

$$D_{\alpha}(x) = e^{v(\alpha)} E D_{\alpha}(x + Y_{\alpha}) - G_{\alpha}(x).$$

Replacing x by x+y in the last expression, dividing both sides by $D_{\alpha}(x)$, and setting $h_{\alpha}(y) = D_{\alpha}(x+y)/D_{\alpha}(x)$ gives

$$h_x(y) = e^{v(\alpha)} E h_x(y + Y_\alpha) - \frac{G_\alpha(x)}{D_\alpha(x)}.$$

If the second term on the right hand side were not there then we would know all the solutions of this equation. To get rid of it we let $x \to \infty$ and show

- (i) $h_x(\cdot)$, $x \in \mathbb{R}$, are uniformly bounded and equicontinuous on compact subsets, so subsequential limits exist,
 - (ii) if h is a subsequential limit then

$$h(y) = e^{v(\alpha)} E h(y + Y_{\alpha})$$

and consequently

(iii) h must be $c_1 e^{(\alpha-\beta_1)x} + c_2 e^{(\alpha-\beta_2)x}$ where β_1 and β_2 are the (at most two) β 's which have $v(\beta) = 0$. That these β 's lie in (0,1] follows from the monotonicity and convexity of φ .

The last statement is possible only if $v(\beta) = 0$ has at least one root in (0, 1] so this completes Case IV. To close the gap between this and the existence result we observe that if α is the smaller root then the convexity of v implies that $v'(\alpha) \leq 0$.

Having found a necessary and sufficient condition for the existence of fixed points the next logical step is to describe the set of all fixed points. As we mentioned before, when there is a fixed point of finite mean there is typically exactly a one parameter family of them, so given the method of constructing fixed points in Cases II and III it seems natural to conjecture that is also the case in general. This turns out to be true if one considers "typically" to mean "under suitable irreducibility assumptions."

We will say that the problem is of lattice type if there is an s>0 so that

with probability one each $\log W_i$ is an integer multiple of s. We will always take s to be the largest possible such number and will refer to it as the span. If the problem is nonlattice we will set s=0. The set of all fixed points is parametrized by a class $\mathfrak{P}_{\alpha,s}$ which we will now define. If s>0 and $\alpha\in(0,1)$, let $\mathfrak{P}_{\alpha,s}$ be the collection of all strictly positive infinitely differentiable functions p on R^1 which satisfy

(a) p(x+s) = p(x) for all $x \in \mathbb{R}^1$, and

(b)
$$(-1)^k \frac{d^k}{d\theta^k} [\theta^{\alpha} p(-\log \theta)] \le 0$$

for all k=1, 2, ... and all $\theta \in (0, \infty)$. As will be seen in Sect. 5, $\mathfrak{P}_{\alpha,s}$ is rather large for these choices of α and s. If s=0 and $\alpha \in (0, 1]$ or if $\alpha=1$ and s>0, let $\mathfrak{P}_{\alpha,s}$ be the set of positive constant functions on R^1 . (It is interesting to note that periodicities play a role if $\alpha < 1$, but not if $\alpha = 1$.)

Note that $v \equiv 0$ if and only if $P(W_i = 1) + P(W_i = 0) = 1$ for each i and $\sum_{i=1}^{N} P(W_i = 1) = 1$, and in this case $\mathfrak{F} = \mathfrak{M} \setminus \{\delta_0\}$. Thus in the next result we will assume that v is not identically zero, in which case there is at most one $\alpha \in (0, 1]$ at which $v(\alpha) = 0$ and $v'(\alpha) \leq 0$. When $\mathfrak{F} \neq \emptyset$, Theorem 1 guarantees that there is exactly one such α .

Theorem 2. Suppose that v is not identically zero and that $\mathfrak{F} \neq \emptyset$. Let α be the unique point in (0,1] for which $v(\alpha)=0$ and $v'(\alpha)\leq 0$.

(a) Then there is a natural one-to-one correspondence between $\varphi \in \mathfrak{F}$ and $p \in \mathfrak{P}_{\sigma}$, which is given by

$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta)} = 1 \quad \text{if } v'(\alpha) < 0,$$

and

$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta) |\log \theta|} = 1 \quad \text{if } v'(\alpha) = 0.$$

(b) Suppose $\varphi \in \mathbb{R}$ and $\psi \in \mathfrak{L}$. If

$$\lim_{\theta \to 0} \frac{1 - \varphi(\theta)}{1 - \psi(\theta)} = 1,$$

then $\lim_{n\to\infty} T^n \psi = \varphi$.

Part (a) says that the fixed points can be identified by the behavior of their Laplace transforms at 0. When $\alpha = 1$ and $v'(\alpha) < 0$ this just says that the fixed points are parametrized by their means. Part (b) is our convergence theorem. It says that if the behavior of ψ matches that of φ at 0 then $T^n\psi \to \varphi$ as $n \to \infty$. When $\alpha = 1$ and $v'(\alpha) < 0$ this just says that the two random variables have the same mean. In other cases the condition is more restrictive, but this situation is analogous to that for sums of independent random variables: when the variance is finite, $(S_n - nm)/n^{1/2}$ converges to a normal, but when we deal with convergence to stable laws we require special assumptions on the tail of the

distribution = information about the behavior of the characteristic function at 0.

Our last result concerns the size of the tails of the distributions of the fixed points in the finite mean case.

Theorem 3. Suppose v(1) = 0 and v'(1) < 0. Then for any $\mu \in \mathcal{F}$ and $\beta > 1$

$$\int_{0}^{\infty} x^{\beta} d\mu < \infty \quad \text{if and only if } v(\beta) < 0.$$

The proof of this is a straightforward generalization of the proof of the corresponding result of Kahane and Peyriere (1976). To see what it says in a concrete case let $W_i = e^{Z_i}$ where Z_i has the normal distribution with mean m and variance $\sigma^2 > 0$. Then $\mathfrak{F} \neq \emptyset$ if and only if

$$m \le -\sigma \sqrt{2 \log N}$$
 if $\sigma > \sqrt{2 \log N}$

and

$$m \le -\log N - \frac{\sigma^2}{2}$$
 if $\sigma \le \sqrt{2\log N}$.

Elements of \mathfrak{F} have a finite moment of order $\beta \ge 1$ if and only if $m = -\log N - \frac{\sigma^2}{2}$ and $\frac{\beta \sigma^2}{2} < \log N$.

It is also interesting to apply this result to the first example considered above. If we let $p < p_c$ and $A = (1 - \sqrt{1 - 16p(1-p)})/4p$ so v(1) = 0, then as $p \to 0$ the number of moments which are finite increases to ∞ . Reversing this and recalling how the fixed points were constructed in Case II we see that as we move up the curve the number of moments which are finite decreases to 0.

In view of the results of Holley and Liggett (1981), our theorems suggest the following conjecture about the behavior of the generalized smoothing process: the existence or nonexistence of stationary measures should correspond to the transience of recurrence of the underlying symmetrized random walk, while the effect of the spread of W should be only to determine the size of the tails of the stationary measure when it exists. If this were the case, the behavior of the generalized smoothing and potlatch processes would be consistent with the behavior of critical cluster fields (see Kallenberg (1977) for example). Resolving these questions even in special cases seems to be quite difficult.

Perhaps it should be mentioned at this point that the connection between the generalized smoothing process and the smoothing transformation is stronger than mere analogy. To see this connection, consider the generalized smoothing process η_t corresponding to the random variable W and the simple

random walk transition probabilities $p(x, y) = \frac{1}{2d}$ for |x - y| = 1 on Z^d . Suppose

the translation invariant initial distribution is a product measure, or more generally, has positive correlations in the sense of Harris (1977). Let $\varphi_t(\theta)$ be the Laplace transform of $\eta_t(0)$. By Theorem 1.1 of Harris (1977), the distribution of the process at time t has positive correlations as well, so

$$\frac{d}{dt} \varphi_t(\theta) = E \exp\left[-\theta \frac{W}{2d} \sum_{|x|=1} \eta_t(x)\right] - \varphi_t(\theta)$$

$$\geq \prod_{|x|=1} E \exp\left[-\theta \frac{W}{2d} \eta_t(x)\right] - \varphi_t(\theta)$$

$$= \left\{E \exp\left[-\theta \frac{W}{2d} \eta_t(0)\right]\right\}^{2d} - \varphi_t(\theta)$$

$$= T\varphi_t(0) - \varphi_t(\theta),$$

where T is the smoothing transformation corresponding to $W_i = \frac{W}{2d}$ for $1 \le i \le 2d$. Therefore

$$\varphi_t(\theta) \geq e^{-t} \sum_{n=0}^{\infty} \frac{t^n}{n!} T^n \varphi_0(\theta),$$

so that $\lim_{n\to\infty} T^n \varphi_0(\theta) = 1$ for all θ implies that $\lim_{n\to\infty} \varphi_t(\theta)$ for all θ .

The proofs of our results are organized according to the tools used. In Sect. 2 the arguments use random walk ideas (potential theory and renewal theory); Section 3 concerns the stable law transformation; Section 4 deals with an associated branching random walk. Section 5 contains the parts of the proofs which do not fit naturally into the earlier sections. Theorem 1 is obtained by combining Theorems 2.7, 2.12, 3.1, 3.2, and 3.5. Theorem 2a is obtained from Theorems 2.18, 4.2, and 5.1. Theorems 2b and 3 are just Theorems 4.2 and 5.3, respectively.

2. The Associated Random Walk

Throughout this section, α will always be a point in (0,1]. Let Y_{α} be a random variable with distribution determined by

(2.1)
$$Ef(Y_{\alpha}) = e^{-v(\alpha)} \sum_{i=1}^{N} E[f(-\log W_i) W_i^{\alpha}, W_i > 0]$$

for nonnegative continuous functions f on R^1 . This is possible since the right hand side of (2.1) is a positive linear functional with unit norm. The random walk to be used in this section is the one whose increments have the distribution of Y_{α} . Initially, the arguments are similar to those used in Holley and Liggett (1981).

Given $\varphi \in \mathcal{L}$ with $\varphi \not\equiv 1$, define $D_{\alpha}(x)$ and $G_{\alpha}(x)$ by

(2.2)
$$D_{\alpha}(x) = e^{\alpha x} [1 - \varphi(e^{-x})], \text{ and}$$

$$G_{\alpha}(x) = e^{\alpha x} E \left\{ \prod_{i=1}^{N} \varphi(e^{-x} W_i) - 1 + \sum_{i=1}^{N} [1 - \varphi(e^{-x} W_i)] \right\}.$$

Put $\tilde{\varphi} = T\varphi$, and define \tilde{D}_{α} and \tilde{G}_{α} analogously in terms of $\tilde{\varphi}$.

(2.3) **Lemma.**
$$\tilde{D}_{\alpha}(x) = e^{v(\alpha)} E D_{\alpha}(x + Y_{\alpha}) - G_{\alpha}(x)$$
.

Proof. Using the appropriate definitions, one checks that

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$$\begin{split} D_{\alpha}(x) &= e^{\alpha x} \left[1 - \tilde{\varphi}(e^{-x}) \right] \\ &= e^{\alpha x} E \left[1 - \prod_{i=1}^{N} \varphi(e^{-x} W_i) \right] \\ &= e^{\alpha x} E \sum_{i=1}^{N} \left[1 - \varphi(e^{-x} W_i) \right] - G_{\alpha}(x) \\ &= \sum_{i=1}^{N} E \left[W_i^{\alpha} D_{\alpha}(x - \log W_i), W_i > 0 \right] - G_{\alpha}(x) \\ &= e^{\varphi(\alpha)} E D_{\alpha}(x + Y_{\alpha}) - G_{\alpha}(x). \end{split}$$

- (2.4) **Lemma.** (a) $G_{\alpha}(x) \ge 0$.
 - (b) $e^{-\alpha x}G_{\alpha}(x)$ is a decreasing function of x.
 - (c) If $\tilde{\varphi} \ge \varphi$, then $\tilde{G}_{\alpha}(x) \le G_{\alpha}(x)$.

Proof. All three statements follow from the following fact: if $0 \le u_i \le v_i \le 1$, then

(2.5)
$$\prod_{i=1}^{N} u_i - 1 + \sum_{i=1}^{N} (1 - u_i) \ge \prod_{i=1}^{N} v_i - 1 + \sum_{i=1}^{N} (1 - v_i).$$

To verify this fact, note that

$$\frac{\partial}{\partial x_i} \left[\prod_{i=1}^N x_i - 1 + \sum_{i=1}^N (1 - x_i) \right] = \prod_{i \neq j} x_i - 1 \le 0.$$

(2.6) **Lemma.** (a) $G_{\alpha}(x) \leq e^{\alpha x} EF(ZD_{\alpha}(x)e^{-\alpha x})$, where $Z = \sum_{i=1}^{N} \max(W_i, 1)$ and

$$F(u) = \begin{cases} e^{-u} - 1 + u & \text{if } u \leq N \\ e^{-N} - 1 + N & \text{if } u \geq N \end{cases}$$

(b)
$$\lim_{x \to \infty} \frac{G_{\alpha}(x)}{D_{\alpha}(x)} = 0.$$

Proof. For part (a), use the inequality $u \le e^{-(1-u)}$ to obtain

$$\begin{aligned} G_{\alpha}(x) & \leq e^{\alpha x} E\left\{e^{-\sum\limits_{i=1}^{N}\left[1-\varphi(e^{-x}W_{i})\right]}-1+\sum\limits_{i=1}^{N}\left[1-\varphi(e^{-x}W_{i})\right]\right\} \\ & \leq e^{\alpha x} EF\left[\sum\limits_{i=1}^{N}\left[1-\varphi(e^{-x}W_{i})\right]\right] \end{aligned}$$

since $\sum_{i=1}^{N} [1 - \varphi(e^{-x}W_i)] \le N$. Since $\varphi \in \mathfrak{Q}$, $\frac{1 - \varphi(u)}{u}$ is decreasing and $1 - \varphi(u)$ is increasing in u. Therefore

$$1 - \varphi(e^{-x}W_i) \leq \max(W_i, 1) [1 - \varphi(e^{-x})].$$

Part (a) now follows from the monotonocity of F on $[0, \infty)$. For part (b), note that since

$$\lim_{x \to \infty} D_{\alpha}(x)e^{-\alpha x} = \lim_{x \to \infty} [1 - \varphi(e^{-x})] = 0,$$

it suffices, by changing variables $t = D_{\alpha}(x)e^{-\alpha x}$, to show that

$$\lim_{t\to\infty}\frac{EF(Zt)}{t}=0.$$

But this follows from the dominated convergence theorem since $\frac{F(u)}{u}$ is bounded and tends to zero as $u \downarrow 0$, and since Z has a finite first moment.

(2.7) **Theorem.** Suppose v(1) = 0 and v'(1) < 0. Then \mathfrak{F} contains elements of finite mean.

Proof. Put $\varphi_0(\theta) = e^{-\theta}$ and define $\varphi_n(\theta)$ recursively by $\varphi_{n+1} = T\varphi_n$. By Jensen's inequality,

$$\varphi_1(\theta) = E \exp \left[-\theta \sum_{i=1}^{N} W_i \right] \ge \exp \left[-\theta E \sum_{i=1}^{N} W_i \right] = \varphi_0(\theta),$$

since v(1) = 0. Therefore

$$(2.8) \varphi_{n+1}(\theta) \ge \varphi_n(\theta)$$

for all n. As a consequence,

$$\varphi(\theta) = \lim_{n \to \infty} \varphi_n(\theta) \ge \varphi_0(\theta)$$

exists and is a fixed point for T. Since $\varphi(\theta) \ge e^{-\theta}$, the measure in $\mathfrak M$ with Laplace transform φ has a mean which is at most one. In order to show that $\varphi \in \mathfrak F$, it then suffices to show that φ is not identically one. In order to do this, let D_n and G_n be defined as in (2.2) in terms of φ_n , with $\alpha = 1$. By (2.8) and Lemma 2.4 (c), $G_{n+1}(x) \le G_n(x)$ for all $n \ge 0$ and $x \in R^1$. Since v(1) = 0, Lemma 2.3 gives

$$D_{n+1}(x) = ED_n(x+Y_1) - G_n(x)$$

$$\geq ED_n(x+Y_1) - G_0(x).$$

Iterating this, we see that

(2.9)
$$D_n(x) \ge ED_0(x+S_n) - E\sum_{k=0}^{n-1} G_0(x+S_k)$$

where S_n is the random walk with $S_0 = 0$ and increments which are distributed like Y_1 . By (2.1),

$$EY_1 = -\sum_{i=1}^{N} EW_i \log W_i = -v'(1) > 0,$$

so that $\lim_{n\to\infty} S_n = +\infty$ a.s. Since $D_0(x)$ is bounded and $\lim_{x\to\infty} D_0(x) = 1$,

(2.10)
$$\lim_{n \to \infty} ED_0(x + S_n) = 1.$$

Since $D_0(x) \le 1$ and F is nondecreasing, Lemma 2.6a gives (recall $\alpha = 1$)

$$\int_{-\infty}^{\infty} G_0(x) dx \le E \int_{-\infty}^{\infty} e^x F(Ze^{-x}) dx$$

$$= E \int_{0}^{\infty} Z \frac{F(\tau)}{\tau^2} d\tau,$$

which is finite since $EZ < \infty$, $F(\tau)$ is bounded, and $F(\tau) \sim \frac{1}{2} \tau^2$ as $\tau \downarrow 0$. By Lemma 2.4b, $e^{-x} G_0(x)$ is decreasing in x. Therefore G_0 is directly Riemann integrable, so by the renewal theorem,

$$\lim_{x \to \infty} E \sum_{k=0}^{\infty} G_0(x + S_k) = 0.$$

This, together with (2.9) and (2.10) implies that

$$\lim_{x\to\infty} \lim_{n\to\infty} D_n(x) = 1.$$

Therefore, by the definition of $D_n(x)$,

$$\lim_{x \to \infty} e^{x} [1 - \varphi(e^{-x})] = 1,$$

so φ is the Laplace transform of a distribution of mean one. This completes the proof of the theorem.

The next theorem involves ideas from the Martin boundary theory for the random walk which corresponds to Y_{α} . In preparation for that we need the following lemma.

(2.11) **Lemma.** Fix $\alpha \in (0, 1]$ and assume that the problem is nonlattice. Let \mathfrak{H}_{α} be the set of all functions g on R^1 which satisfy (i) g(0)=1, (ii) $g(y)e^{-\alpha y}$ is decreasing in y, (iii) $g(y)e^{(1-\alpha)y}$ is increasing in y, and (iv)

$$g(y) = e^{v(\alpha)} E g(y + Y_{\alpha}).$$

Then \mathfrak{H}_{α} is the set of convex combinations of $g_{\beta}(y) = e^{(\alpha - \beta)y}$ for the (at most two) β 's which satisfy $0 \le \beta \le 1$ and $v(\beta) = 0$.

Proof. It is easy to check that $g_{\beta} \in \mathfrak{H}_{\alpha}$ if and only if $0 \le \beta \le 1$ and $v(\beta) = 0$. Therefore the statement of the lemma is immediate if $\mathfrak{H}_{\alpha} = \emptyset$. Assume then that \mathfrak{H}_{α} is not empty. \mathfrak{H}_{α} is a compact convex subset of $C(R^1)$ with the topology of uniform convergence on bounded sets. Hence it is the closed convex hull of its extreme points by the Krein-Milman theorem (see Royden (1968), for example). Suppose that g is an extreme point of \mathfrak{H}_{α} and let

$$g_u(y) = \frac{g(u+y)}{g(u)}$$
.

By property (iv) of the definition of \mathfrak{H}_{α} and the fact that $Eg(Y_{\alpha}) = e^{-v(\alpha)}$,

$$g(y) = \frac{Eg(y + Y_{\alpha})}{Eg(Y_{\alpha})} = \frac{\int_{-\infty}^{\infty} g_{u}(y)g(u)P[Y_{\alpha} \in du]}{\int_{-\infty}^{\infty} g(u)P[Y_{\alpha} \in du]}.$$

Since $g_u \in \mathfrak{H}_{\alpha}$ for each u and g is extremal, this implies that $g = g_u$ for all u in the support of the distribution of Y_a . Therefore

$$g(u+y)=g(u)g(y)$$

for all y and all u in the support of Y_{α} . Since g is continuous and the problem is nonlattice, it follows that $g = g_{\beta}$ for some β , thus completing the proof of the lemma.

(2.12) **Theorem.** Suppose that $\varphi \in \mathfrak{F}$. Then (a) there is an $\alpha \in [0,1]$ so that $v(\alpha) = 0$, and (b) if v is not identically zero and $\alpha \in (0,1]$ is such that $v(\alpha) = 0$ and $v'(\alpha) \leq 0$, then

$$\limsup_{x \to \infty} \frac{D_{\alpha}(x+y)}{D_{\alpha}(x)} \le 1 \quad \text{if } v'(\alpha) < 0$$

and

$$\lim_{x \to \infty} \frac{D_{\alpha}(x+y)}{D_{\alpha}(x)} = 1 \quad \text{if } v'(\alpha) = 0$$

where y>0 is any multiple of s if the problem is of lattice type, and is arbitrary otherwise.

Proof. Fix an $\alpha \in (0, 1]$, and put

$$h_x(y) = \frac{D_\alpha(x+y)}{D_\alpha(x)}.$$

By Lemma 2.3,

$$D_{\alpha}(x) = e^{v(\alpha)} E D_{\alpha}(x + Y_{\alpha}) - G_{\alpha}(x).$$

Evaluating this at (x + y) and dividing by $D_{\alpha}(x)$ gives

(2.13)
$$h_x(y) = e^{v(\alpha)} E h_x(y + Y_\alpha) - \frac{G_\alpha(x+y)}{D_\alpha(x+y)} h_x(y).$$

Since $\varphi \in \mathfrak{L}$, $D_{\alpha}(x)e^{-\alpha x} = 1 - \varphi(e^{-x})$ is decreasing in x and $D_{\alpha}(x)e^{(1-\alpha)x}$ is increasing in x. Therefore

$$(2.14) -(1-\alpha)D_{\alpha}(x) \leq D'_{\alpha}(x) \leq \alpha D_{\alpha}(x),$$

so evaluating this at (x + y) and dividing by $D_{\alpha}(x)$ gives

$$-(1-\alpha)h_x(y)\!\leq\!h_x'(y)\!\leq\!\alpha\,h_x(y).$$

Since $h_x(0) = 1$, it follows that

$$h_x(y) \leq \begin{cases} e^{\alpha y} & \text{for } y > 0 \\ e^{-(1-\alpha)y} & \text{for } y < 0, \end{cases}$$

which by the previous inequalities gives bounds on $h'_x(y)$. In particular, the collection $\{h_x(\cdot), x \in R^1\}$ is uniformly bounded and equicontinuous on bounded subsets of R^1 , and hence is relatively compact in the topology of uniform convergence on bounded sets. By the definition of Y_α ,

(2.15)
$$E(e^{\alpha Y_{\alpha}}) = e^{v(0) - v(\alpha)} < \infty, \quad \text{and} \quad E(e^{-(1 - \alpha)Y_{\alpha}}) = e^{v(1) - v(\alpha)} < \infty.$$

Now suppose $x_n \to \infty$ and $h_{x_n}(y) \to h(y)$ uniformly on bounded sets of R^1 . Then by Lemma 2.6b and the dominated convergence theorem, we may pass to the limit in (2.13) to obtain

$$(2.16) h(y) = e^{v(\alpha)} E h(y + Y_{\alpha}).$$

Assume from now on that the problem is nonlattice. The lattice case is similar. Then $h \in \mathfrak{H}_{\alpha}$, so by Lemma 2.11, there is a $\beta \in [0,1]$ for which $v(\beta) = 0$. This proves part (a) of the theorem. For part (b), suppose now that $v(\alpha) = 0$ and $v'(\alpha) \leq 0$. Since we are in the nonlattice case, v is not identically zero. By the convexity of v, if $v(\beta) = 0$ then $\beta \geq \alpha$ if $v'(\alpha) < 0$ and $\beta = \alpha$ if $v'(\alpha) = 0$. Therefore by Lemma 2.11, h(y) is decreasing in v if $v'(\alpha) < 0$ and is constant if $v'(\alpha) = 0$. Since $v(\alpha) = 0$ if $v(\alpha) = 0$ since this is true for all limit points of $v(\alpha) = 0$ and $v(\alpha) = 0$ in the proof of the theorem is complete.

(2.17) Corollary. Suppose that $\varphi \in \mathfrak{F}$, $\alpha \in (0,1]$, $v(\alpha) = 0$, $v'(\alpha) \leq 0$, and v is not identically zero. Then $G_{\alpha}(x)$ is directly Riemann integrable on R^1 . (See p. 348 of Feller (1966) for the definition.)

Proof. By Lemma 2.4b, $e^{-\alpha x}G_{\alpha}(x)$ is decreasing in x. The first step is to use this monotonicity to show that if G_{α} is integrable, then it is directly Riemann integrable. To do this, take h>0, and let $\underline{m}_n(h)$ and $\overline{m}_n(h)$ be the minimum and maximum values of G_{α} on the interval [(n-1)h, nh]. Then by the monotonicity of $e^{-\alpha x}G_{\alpha}(x)$,

$$\bar{m}_n(h) \leq e^{\alpha h} G_{\alpha}[(n-1)h]$$

and

$$\underline{m}_n(h) \geq e^{-\alpha h} G_{\alpha}[nh].$$

Therefore

$$\begin{split} \sum_{n=L}^{M} \left[\overline{m}_n(h) - \underline{m}_n(h) \right] & \leq e^{\alpha h} \sum_{n=L-1}^{M-1} G_{\alpha} [nh] - e^{-\alpha h} \sum_{n=L}^{M} G_{\alpha} [nh] \\ & \leq \left[e^{\alpha h} - e^{-\alpha h} \right] \sum_{n=-\infty}^{\infty} G_{\alpha} [nh] + e^{\alpha h} G_{\alpha} ((L-1)h). \end{split}$$

Using the monotonicity of $e^{-\alpha} G_{\alpha}(x)$ again,

$$\int_{(n-1)h}^{nh} G_{\alpha}(x) \ge h e^{-\alpha h} G_{\alpha}(nh).$$

Therefore if $\int_{-\infty}^{\infty} G_{\alpha}(x)dx < \infty$, it follows that $\sum_{n} \bar{m}_{n}(h) < \infty$ and $\sum_{n} \underline{m}(h) < \infty$, and

that

$$h \sum_{n=-\infty}^{\infty} \left[\overline{m}_n(h) - \underline{m}_n(h) \right] \leq (e^{2\alpha h} - 1) \int_{-\infty}^{\infty} G_{\alpha}(x) dx.$$

Since the right side of this inequality tends to zero as $h \downarrow 0$, G_{α} is directly Riemann integrable. To show that G_{α} is integrable, note that by Lemmas 2.4a and 2.6a,

$$0 \le G_{\alpha}(x) \le e^{\alpha x} EF(ZD_{\alpha}(x)e^{-\alpha x}).$$

Since F is bounded,

$$\int_{-\infty} G_{\alpha}(x) dx < \infty.$$

To deal with integrability at $+\infty$, choose $\beta > 0$ so small that $\beta < \alpha/2$ and $\frac{\alpha}{\alpha - \beta} \le \gamma$. Since $D_{\alpha}(x)e^{-\alpha x}$ is decreasing in x, Theorem 2.12b implies that

$$\limsup_{x \to \infty} \frac{1}{x} \log D_{\alpha}(x) \leq 0.$$

Therefore there is an x_0 so that for $x \ge x_0$

$$D_{\alpha}(x) \leq e^{\beta x}$$
.

Since F is monotone,

$$G_{\alpha}(x) \leq e^{\alpha x} EF(Ze^{(\beta-\alpha)x})$$

for $x \ge x_0$. By making a change of variable $u = Ze^{(\beta - \alpha)x}$ in the integral and using Fubini's theorem (the integrand is nonnegative),

$$\int_{x_0}^{\infty} G_{\alpha}(x) dx \leq \int_{x_0}^{\infty} e^{\alpha x} EF(Ze^{(\beta - \alpha)x}) dx$$

$$= E \int_{x_0}^{\infty} e^{\alpha x} F(Ze^{(\beta - \alpha)x}) dx$$

$$\leq E \int_{0}^{\infty} \frac{Z^{\alpha/(\alpha - \beta)}}{\alpha - \beta} \frac{F(u)}{u^{1 + (\alpha/(\alpha - \beta))}} du.$$

The last integral is finite since $EZ^{\alpha/(\alpha-\beta)} < \infty$, F is bounded, and $F(u) \sim u^2/2$ as $u \downarrow 0$.

The following result is the key to identifying the elements of §.

(2.18) **Theorem.** Suppose that v is not identically zero and that $\alpha \in (0, 1]$ satisfies $v(\alpha) = 0$ and $v'(\alpha) \leq 0$. If $\varphi \in \mathfrak{F}$ then there is a $p \in \mathfrak{P}_{\alpha, s}$ so that

(a)
$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta)} = 1$$
 if $v'(\alpha) < 0$, and

(b)
$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} |\log \theta| \, p(-\log \theta)} = 1 \text{ if } v'(\alpha) = 0.$$

Proof. Let S_n be the random walk with $S_0 = 0$ whose increments have distribution Y_α . Since $v(\alpha) = 0$ and $\varphi \in \mathfrak{F}$, Lemma 2.3 gives

(2.19)
$$D_{\alpha}(x) = ED_{\alpha}(x + Y_{\alpha}) - G_{\alpha}(x).$$

This is known as Poisson's equation for the random walk S_n . In order to prove our theorem, it is necessary to determine the behavior at $+\infty$ of the solution $D_{\alpha}(x)$ of this equation. The result we need would follow from the general theory of Poisson's equation if either the random walk were transient, or if it were recurrent and its distribution were either lattice or nonsingular (see Port and Stone (1969), for example). In our situation, however, these assumptions are not necessary for the desired conclusion, so we will give our own derivation of the necessary facts which takes advantage of special properties of our solution $D_{\alpha}(x)$. We begin with the case $v'(\alpha) < 0$. In this case S_n is transient since

$$E Y_{\alpha} = -\sum_{i=1}^{N} E W_{i}^{\alpha} \log W_{i} = -v'(\alpha) > 0.$$

Iterating (2.19) and passing to the limit, we see that

$$D_{\alpha}(x) = \lim_{n \to \infty} ED_{\alpha}(x + S_n) - \sum_{k=0}^{\infty} EG_{\alpha}(x + S_k).$$

Here the sum is finite by Corollary 2.17, the renewal theorem (Theorem 1 of Sect. XI.9 of Feller (1966)), and the fact that $EY_{\alpha} > 0$, while the limit exists because $ED_{\alpha}(x+S_n)$ is increasing in n. Of course

$$p(x) = \lim_{n \to \infty} ED_{\alpha}(x + S_n) \ge D_{\alpha}(x) > 0$$

is harmonic for the random walk. Since $D_{\alpha}(x)e^{-\alpha x}$ is decreasing in x and $D_{\alpha}(x)e^{(1-\alpha)x}$ is increasing in x, p(x) is continuous. By the renewal theorem and Corollary 2.17,

$$\sum_{k=0}^{\infty} EG_{\alpha}(x+S_k)$$

is bounded on R^1 and tends to zero as $x \to +\infty$. Since $D_{\alpha}(x) \le e^{\alpha x}$, $\overline{\lim}_{x \to -\infty} p(x) < \infty$.

Therefore since $EY_{\alpha} > 0$, p(x) is constant in the nonlattice case and periodic of period s in the lattice case. Hence

$$\lim_{x \to \infty} [D(\alpha) - p(x)] = -\lim_{x \to \infty} \sum_{k=0}^{\infty} EG_{\alpha}(x + S_k) = 0.$$

Putting $\theta = e^{-x}$ and recalling the definition of D_{α} , it follows that

(2.20)
$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta)} = 1.$$

To check that in the lattice case

$$(-1)^k \frac{d^k}{d\theta^k} [\theta^{\alpha} p(-\log \theta)] \leq 0$$

for all $k \ge 1$ and $\theta > 0$, use the periodicity of p to write

$$\theta^{\alpha} p(-\log \theta) = \theta^{\alpha} p(-\log \theta + ns)$$

$$= \theta^{n\alpha s} \left[\frac{\theta^{\alpha} e^{-n\alpha s} p(-\log \theta + ns)}{1 - \varphi(\theta e^{-ns})} \right] [1 - \varphi(\theta e^{-ns})].$$

Therefore

$$\theta^{\alpha} p(-\log \theta) = \lim_{n \to \infty} e^{n\alpha s} [1 - \varphi(\theta e^{-ns})].$$

Since $\varphi \in \mathfrak{Q}$, it follows that the derivatives of $\theta^{\alpha} \rho(-\log \theta)$ have the correct signs. Note that since $\frac{1-\varphi(\theta)}{\theta}$ is monotone, if $\alpha=1$ then p is both monotone and periodic, and hence constant. Turning now to the recurrent case, assume that $v'(\alpha)=0$. Let τ be the first time that S_n enters $(0,\infty)$, so that S_{τ} is the strict ascending ladder variable associated with Y_{α} . Recall that $Y_{\alpha} \neq 0$ since $v \neq 0$ and that $EY_{\alpha}=0$ since $v'(\alpha)=0$. Therefore $\tau<\infty$ a.s. By (2.19),

$$D_{\alpha}(x+S_n) - \sum_{k=0}^{n-1} G_{\alpha}(x+S_k)$$

is a martingale. By the martingale stopping theorem,

(2.21)
$$ED_{\alpha}(x + S_{\tau \wedge n}) - E \sum_{k=0}^{\tau \wedge n-1} G_{\alpha}(x + S_k) = D_{\alpha}(x).$$

By (2.15) and (3.6a) of Chap. XII of Feller (1966),

$$(2.22) Ee^{\alpha S_{\tau}} < \infty.$$

Therefore, since $D_{\alpha}(x) \leq e^{\alpha x}$ and $S_n \leq S_{\tau}$ for $n < \tau$, we may pass to the limit in (2.21) to obtain

$$ED_{\alpha}(x+S_{\tau})-E\sum_{k=0}^{\tau-1}G(x+S_{k})=D_{\alpha}(x).$$

This is again Poisson's equation, but for the random walk whose increments have the distribution of S_{τ} . Put

(2.23)
$$R(x) = E \sum_{k=0}^{\tau-1} G_{\alpha}(x+S_k) = ED_{\alpha}(x+S_{\tau}) - D_{\alpha}(x).$$

By the duality lemma of Sect. XII.2 of Feller (1966),

$$R(x) = \sum_{k=0}^{\infty} EG_{\alpha}(x + T_k)$$

where T_k is the random walk whose increments have the distribution of the weak descending ladder variable for the original random walk S_n . So, by the

renewal theorem and Corollary 2.17, there is a strictly positive continuous function p(x) which is constant in the nonlattice case and periodic of period s in the lattice case so that

(2.24)
$$\lim_{x \to \infty} [R(x) - p(x)] = 0.$$

Consider now the nonlattice case only, since the lattice case is handled similarly with derivatives and integrals being replaced by differences and sums respectively. Rewrite (2.23) as

(2.25)
$$E \int_{0}^{S_{\tau}} D'_{\alpha}(z+y) dy = R(z).$$

Since $D_{\alpha}(x) \leq e^{\alpha x}$, (2.14) and (2.22) can be used to justify the interchange of expectation and integration in (2.25) to obtain

$$\int_{0}^{\infty} D'_{\alpha}(z+y)P(S_{\tau} \geq y)dy = R(z).$$

Integrating both sides of this identity with respect to z (using (2.14) and (2.22) again to justify the interchange of order of integration),

(2.26)
$$\int_{0}^{\infty} D_{\alpha}(x+y)P(S_{\tau} \ge y)dy = \int_{0}^{x} R(z)dz + c$$

where $c = \int_{0}^{\infty} D_{\alpha}(y) P(S_{\tau} \ge y) < \infty$. Since $D_{\alpha}(x) e^{-\alpha x}$ is decreasing in x,

$$\frac{D_{\alpha}(x+y)}{D_{\alpha}(x)} \leq e^{\alpha y}$$

for $y \ge 0$. Therefore, dividing (2.26) by $D_{\alpha}(x)$ and using (2.22), Theorem 2.12b, and the dominated convergence theorem, we see that

$$\lim_{x\to\infty} \frac{\int\limits_0^\infty R(z)dz + c}{D_{\alpha}(x)} = ES_{\tau},$$

which is positive and finite. Therefore, by (2.24),

$$\lim_{x \to \infty} \frac{D_{\alpha}(x)}{x} = \frac{1}{ES_{\tau}} \lim_{x \to \infty} \frac{1}{x} \int_{0}^{x} p(z) dz$$

exists and is positive and finite. In the lattice case, the corresponding conclusion is that

(2.27)
$$\lim_{n\to\infty} \frac{D_{\alpha}(x+ns)}{x+ns} = \frac{p(x)}{ES_{\tau}}$$

for each $x \in \mathbb{R}^{1}$. To deduce from this that

(2.28)
$$\lim_{x \to \infty} \frac{D_{\alpha}(x)}{x p(x)} = \frac{1}{ES_{\tau}},$$

proceed as follows. Let

$$u_n(x) = \frac{D_{\alpha}(x + ns)}{x + ns}$$

for $0 \le x \le s$. Then (2.27) asserts the pointwise convergence of u_n , while (2.28) asserts the uniform convergence of u_n on [0, s]. But by the monotonicity of $D_{\alpha}(x)e^{-\alpha x}$ and (2.27), $D_{\alpha}(x)/x$ is bounded at $+\infty$. Hence by (2.14), $D_{\alpha}(x)/x$ has a uniformly bounded derivative at $+\infty$. Hence the family $\{u_n(x)\}$ is equicontinuous on [0, s], so pointwise convergence implies uniform convergence. Part (b) of the theorem now follows from (2.28) with p(x) replaced by $p(x)/ES_{\tau}$, by putting $\theta = e^{-x}$. The verification that $p(x) \in \mathfrak{P}_{\alpha,s}$ is the same as in the transient case, which was dealt with earlier in this proof.

3. The Stable Transformation

This section is devoted primarily to the construction of fixed points for T of infinite mean. This is done by stopping a one sided stable process at a random time whose distribution is a fixed point of finite mean for a closely related smoothing transformation. This approach is natural since fixed points of finite mean have already been constructed under appropriate assumptions in Theorem 2.7, and since in view of Theorem 2.18, any fixed point must have stable like tails.

(3.1) **Theorem.** Suppose $v(\alpha) = 0$ and $v'(\alpha) < 0$ for some $\alpha \in (0, 1)$. Then \mathfrak{F} is nonempty.

Proof. Define $\tilde{W}_i = W_i^{\alpha}$, and let \tilde{v} , \tilde{T} and $\tilde{\mathfrak{F}}$ be defined in terms of $\{\tilde{W}_1, \ldots, \tilde{W}_N\}$ just as v, T and $\tilde{\mathfrak{F}}$ were defined in terms of $\{W_1, \ldots, W_N\}$. Then $\tilde{v}(\beta) = v(\alpha\beta)$, so that $\tilde{v}(1) = v(\alpha) = 0$ and $\tilde{v}'(1) = \alpha v'(\alpha) < 0$. Therefore by Theorem 2.7, there is a $\psi \in \tilde{\mathfrak{F}}$ with finite mean. Put

$$\varphi(\theta) = \psi(\theta^{\alpha}).$$

Then

$$T\varphi(\theta) = E \prod_{i=1}^{N} \varphi(\theta W_i)$$

$$= E \prod_{i=1}^{N} \psi(\theta^{\alpha} W_i^{\alpha})$$

$$= \tilde{T}\psi(\theta^{\alpha})$$

$$= \psi(\theta^{\alpha}) = \varphi(\theta).$$

To check that $\varphi \in \mathfrak{F}$, it then suffices to verify that φ is a Laplace transform. To see this, let X(t) be the one sided stable process of index α whose Laplace

transform is

$$e^{-t\theta^{\alpha}}$$

(see Sect. XIII.6 of Feller (1966), for example), and let τ be a random variable with Laplace transform ψ which is independent of $X(\cdot)$. Then $X(\tau)$ has Laplace transform φ , since

$$E(e^{-\theta X(\tau)})E[E(e^{-\theta X(\tau)}|\tau)]$$

$$=E[e^{-\theta^{\alpha}\tau}] = \psi(\theta^{\alpha}) = \varphi(\theta).$$

- (3.2) **Theorem.** Suppose that \mathfrak{F} is nonempty and that v is not identically zero. Then
 - (a) v(0) > 0, and
- (b) for any $\phi \in \mathfrak{F}$, $\phi(\infty)$ is equal to the unique fixed point in [0,1) of the function

$$f(t) = \sum_{k=0}^{N} t^{k} P \left[\sum_{i=1}^{N} 1_{\{W_{i} > 0\}} = k \right].$$

Proof. For $\varphi \in \mathfrak{F}$,

(3.3)
$$\varphi(\theta) = E \prod_{i=1}^{N} \varphi(\theta W_i).$$

Therefore, taking the limit as $\theta \uparrow \infty$, we see that

$$\varphi(\infty) = E \prod_{i=1}^{N} \left[1_{\{W_i = 0\}} + \varphi(\infty) 1_{\{W_i > 0\}} \right]$$
$$= f \left[\varphi(\infty) \right],$$

so f has a fixed point in the interval [0, 1). Since f is convex, f(1) = 1, and

$$f'(1) = \sum_{k=0}^{N} k P \left[\sum_{i=1}^{N} 1_{\{W_i > 0\}} = k \right]$$
$$= \sum_{i=1}^{N} P(W_i > 0) = e^{v(0)},$$

it follows that v(0) > 0 or that f(t) = t for all t. In the latter case,

$$\sum_{i=1}^{N} 1_{\{W_i > 0\}} = 1 \text{ a.s.,}$$

so if we let $W = W_1 + ... + W_N$, (3.3) becomes

(3.4)
$$\varphi(\theta) = E \varphi(\theta W),$$

or in terms of random variables, X = XW. Taking logarithms and iterating, we see that this can only happen if $W \equiv 1$, in which case $W_i = 1$ on $\{W_i > 0\}$. This implies that v is identically zero, which is ruled out by the assumptions of the

theorem. Thus v(0) > 0, and f has a unique fixed point in [0, 1), which proves both parts of the theorem.

(3.5) **Theorem.** Suppose that v is not identically zero, but that $v(\alpha) = 0$ and $v'(\alpha) = 0$ for some $\alpha \in (0, 1]$. Then \mathfrak{F} is not empty.

Proof. Since v is strictly convex under the given assumptions, $v(\beta) > 0$ and $v'(\beta) < 0$ for all $\beta \in (0, \alpha)$. Define

$$W_{i,\beta} = W_i e^{-v(\beta)/\beta}$$

for $1 \le i \le N$ and $\beta \in (0, \alpha)$, and let v_{β} and T_{β} be the corresponding function and smoothing transformation. Then

$$\begin{split} v_{\beta}(\delta) &= \log \left[E \sum_{i=1}^{N} W_{i,\beta}^{\delta} \right] \\ &= \log \left[e^{-\delta v(\beta)/\beta} E \sum_{i=1}^{N} W_{i}^{\delta} \right] \\ &= v(\delta) - \frac{\delta v(\beta)}{\beta}. \end{split}$$

Therefore $v_{\beta}(\beta) = 0$ and $v_{\beta}'(\beta) = v'(\beta) - \frac{v(\beta)}{\beta} < 0$ for $\beta \in (0, \alpha)$. Therefore T_{β} has a nontrivial fixed point ψ_{β} by Theorem 3.1. Note that the function f of Theorem 3.2 does not depend on β , so by that theorem, $\psi_{\beta}(\infty) = t_0 \in [0, 1)$ is independent of β . Since $\psi_{\beta}(c\theta)$ is a fixed point of T_{β} whenever $\psi_{\beta}(\theta)$ is, we may choose the fixed point ψ_{β} so that

$$\psi_{\beta}(1) = \frac{t_0 + 1}{2}.$$

Let X_{β} be a random variable with Laplace transform ψ_{β} , and choose a sequence $\beta_n \uparrow \alpha$ so that X_{β_n} converges vaguely to $X \leq \infty$ as $n \to \infty$. Put $\varphi(0) = 1$ and

$$\varphi(\theta) = E(e^{-\theta X}, X < \infty)$$

for $\theta > 0$. Then $\psi_{\beta_n}(\theta) \to \varphi(\theta)$ for all θ , so in particular $\varphi(1) = \frac{1+t_0}{2}$. Since ψ_{β} is a fixed point for T_{β} ,

$$\psi_{\beta}(\theta) = E \prod_{i=1}^{N} \psi_{\beta}(\theta W_{i} e^{-v(\beta)/\beta}).$$

Since $v(\beta_n) \to 0$, we may pass to the limit as $n \to \infty$ to obtain

(3.6)
$$\varphi(\theta) = E \prod_{i=1}^{N} \varphi(\theta W_i).$$

Letting $\theta \downarrow 0$, we see that $\phi(0+)=f[\phi(0+)]$, where f is the function defined in Theorem 3.2. Since

$$\varphi(0+) \geq \varphi(1) > t_0$$

it follows that $\varphi(0+)=1$. Therefore $X<\infty$ a.s., so $\varphi\in\mathfrak{L}$ and hence $\varphi\in\mathfrak{F}$ by (3.6) and the fact that $\varphi(1)<1$.

4. The Associated Branching Random Walk

This section is devoted to the proof of the convergence theorem for iterates of T. This convergence theorem also plays a role in the characterization of \mathcal{F} given in Theorem 2a from the introduction. It is convenient to introduce the following discrete time branching random walk on $[-\infty, \infty)$. At each unit of time, each individual produces N offspring and then dies. If x is the position of the parent, then the offspring are placed at the positions $\{x + \log W_1, \ldots, x + \log W_N\}$, where the vector (W_1, \ldots, W_N) is chosen independently by each parent. Note that this is well defined even if x or $\log W_i$ is $-\infty$. Let η_n be the configuration of the branching random walk at time n. Since

$$E^{\eta} \prod_{\eta_1(x) \geq 1} \left[\varphi(e^x) \right]^{\eta_1(x)} = \prod_{\eta(x) \geq 1} \left[T \varphi(e^x) \right]^{\eta(x)},$$

a simple iteration yields

$$T^n \varphi(\theta) = E^{\{\log \theta\}} \prod_{\eta_n(x) \geq 1} \left[\varphi(e^x) \right]^{\eta_n(x)}.$$

Let $L_n = \max\{x : \eta_n(x) \ge 1\}$ be the position of the rightmost particle.

(4.1) **Lemma.** Suppose that v is not identically zero, but that $v(\alpha) = 0$ for some $\alpha > 0$. Then

$$\lim_{n\to\infty} L_n = -\infty \quad \text{a.s.}$$

Proof. Since v is not identically zero and $v(\alpha) = 0$ for some $\alpha > 0$,

$$P(\max_{1 \le i \le N} W_i = 1) < 1.$$

Therefore, if $P(\max_{\substack{1 \le i \le N \\ 1 \le i \le N}} W_i \le 1) = 1$, $L_n \to -\infty$ a.s. trivially. On the other hand, if $P(\max_{\substack{1 \le i \le N }} W_i > 1) > 0$,

$$P(\limsup_{n\to\infty}L_n=-\infty)+P(\limsup_{n\to\infty}L_n=+\infty)=1.$$

Therefore it suffices to show that $\sup L_n < \infty$ a.s. In order to do this, let $\alpha > 0$ be such that $v(\alpha) = 0$. Then

$$M_n = \sum_{x: \, \eta_n(x) \ge 1} \, \eta_n(x) e^{\alpha x}$$

is a (nonnegative) martingale, since

$$E\sum_{i=1}^{N} e^{\alpha \log W_i} = E\sum_{i=1}^{N} W_i^{\alpha} = 1.$$

This martingale was used by Kahane and Peyriere (1976) and by Kingman (1975). By the martingale convergence theorem, $\lim_{n\to\infty} M_n$ exists and is finite a.s. Since $e^{\alpha L_n} \leq M_n$, it follows that $\sup_n L_n < \infty$ a.s., thus completing the proof of the lemma.

(4.2) **Theorem.** Assume that v is not identically zero. Suppose that $\varphi \in \mathfrak{F}$ and $\psi \in \mathfrak{L}$. If

 $\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{1 - \psi(\theta)} = 1,$

then

$$\lim_{n\to\infty} T^n \psi = \varphi.$$

Proof. By Theorems 3.2a and 2.12a, v(0) > 0 and $v(\alpha) = 0$ for some $\alpha \in (0, 1]$. Let α be the smallest such value. Then $v'(\alpha) \leq 0$. By Theorem 2.18,

 $\liminf_{\theta \downarrow 0} \frac{1 - \varphi(c\theta)}{1 - \varphi(\theta)} = c^{\alpha} \min_{0 \le x \le s} \frac{p(-\log c + x)}{p(x)}$

and

$$\limsup_{\theta \downarrow 0} \frac{1 - \varphi(c\theta)}{1 - \varphi(\theta)} = c^{\alpha} \max_{0 \le x \le s} \frac{p(-\log c + x)}{p(x)}$$

for c>0, where $p \in \mathfrak{P}_{\alpha,s}$. Since $p \in \mathfrak{P}_{\alpha,s}$, $\theta^{\alpha} p(-\log \theta)$ is strictly increasing on $[0,\infty)$. Therefore

$$\liminf_{\theta \downarrow 0} \frac{1 - \varphi(c\theta)}{1 - \varphi(\theta)} > 1 \quad \text{if } c > 1$$

and

$$\limsup_{\theta \downarrow 0} \frac{1 - \varphi(c\theta)}{1 - \varphi(\theta)} < 1 \quad \text{if } c < 1.$$

Fix c > 1 and put

$$\varphi(\theta) = \varphi(c\theta)$$
 and $\bar{\varphi}(\theta) = \varphi(c^{-1}\theta)$.

Then φ , $\bar{\varphi} \in \mathfrak{F}$ and for some $\theta_0 > 0$ and all $0 < \theta \leq \theta_0$,

$$\varphi(\theta) \leq \psi(\theta) \leq \bar{\varphi}(\theta)$$
.

Therefore

$$\prod_{\eta_n(x)\geq 1} \left[\varphi(e^x) \right]^{\eta_n(x)} \leq \prod_{\eta_n(x)\geq 1} \left[\psi(e^x) \right]^{\eta_n(x)} \leq \prod_{\eta_n(x)\geq 1} \left[\bar{\varphi}(e^x) \right]^{\eta_n(x)}$$

on the event $\{L_n \leq \log \theta_0\}$. Since $\bar{\varphi}(\theta) \leq 1$, Lemma 4.1 now gives

$$\lim_{n \to \infty} \inf T^n \underline{\varphi}(\theta) \leq \lim_{n \to \infty} \inf T^n \psi(\theta)
\leq \lim_{n \to \infty} \sup T^n \psi(\theta)
\leq \lim_{n \to \infty} \sup T^n \overline{\varphi}(\theta).$$

Since $T^n \underline{\varphi} = \underline{\varphi}$ and $T^n \overline{\varphi} = \overline{\varphi}$, it follows that all limit points of $T^n \psi(\theta)$ lie between $\varphi(c\theta)$ and $\varphi(c^{-1}\theta)$. Since c > 1 is arbitrary,

$$\lim_{n\to\infty} T^n \psi(\theta) = \varphi(\theta).$$

5. The Totality of Fixed points; Moments

From Theorems 2.7, 3.1 and 3.5, we now know that if $v(\alpha)=0$ and $v'(\alpha)\leq 0$ for some $\alpha\in(0,1]$, then \mathfrak{F} is nonempty. Furthermore, Theorem 2.18 limits the possible behavior at the origin of any $\varphi\in\mathfrak{F}$. The first result in this section says that each behavior permitted by Theorem 2.18 in fact can occur.

(5.1) **Theorem.** Suppose that v is not identically zero and that $v(\alpha)=0$ and $v'(\alpha) \leq 0$ for some $\alpha \in (0,1]$. If $p \in \mathfrak{P}_{\alpha,s}$, then there is a unique $\phi \in \mathfrak{F}$ so that

$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta)} = 1 \quad \text{if } v'(\alpha) < 0$$

and

$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta) |\log \theta|} = 1 \quad \text{if } v'(\alpha) = 0.$$

Proof. The uniqueness comes from Theorem 4.2. If the problem is nonlattice or $\alpha = 1$, then $\mathfrak{P}_{\alpha,s}$ consists only of constants, so this result follows from Theorems 2.7, 2.18, 3.1 and 3.5. So, we can assume that $\alpha < 1$ and s > 0. Let $g(\theta) = e^{-\theta}$ if $v'(\alpha) < 0$ and

$$g(\theta) = \frac{2}{\pi} \int_{0}^{\infty} \frac{e^{-\theta x}}{1 + x^2} dx,$$

which is asymptotic to $1-\theta|\log\theta|$ as $\theta\downarrow 0$, if $v'(\alpha)=0$. Then by criterion 2 of Sect. XIII.4 of Feller (1966), since $p\in\mathfrak{P}_{\alpha,s}$ it follows that

$$\psi(\theta) = g\left[\frac{\theta^{\alpha}p(-\log\theta)}{\alpha}\right]$$

is in \mathfrak{L} . It is easy to check that

$$\lim_{\theta \downarrow 0} \frac{1 - \psi(\theta)}{\theta^{\alpha} p(-\log \theta)} = 1 \quad \text{if } v'(\alpha) < 0$$

and

$$\lim_{\theta \downarrow 0} \frac{1 - \psi(\theta)}{\theta^{\alpha} p(-\log \theta) |\log \theta|} = 1 \quad \text{if } v'(\alpha) = 0.$$

By Theorems 3.1 and 3.5, \mathfrak{F} is nonempty. Take $\tilde{\psi} \in \mathfrak{F}$, which by Theorem 2.18 satisfies

$$\lim_{\theta \downarrow 0} \frac{1 - \tilde{\psi}(\theta)}{\theta^{\alpha} \tilde{p}(-\log \theta)} = 1 \quad \text{if } v'(\alpha) < 0$$

and

$$\lim_{\theta \downarrow 0} \frac{1 - \tilde{\psi}(\theta)}{\theta^{\alpha} \tilde{p}(-\log \theta) |\log \theta|} = 1 \quad \text{if } v'(\alpha) = 0$$

for some $\tilde{p} \in \mathfrak{P}_{\alpha,s}$. Since $\tilde{p} \in \mathfrak{P}_{\alpha,s}$, $\theta^{\alpha} \tilde{p}(-\log \theta)$ is strictly increasing on $[0, \infty)$ and is 0 at 0 and tends to ∞ at ∞ . Hence $u(\theta)$ can be defined by

$$[u(\theta)]^{\alpha} \tilde{p} [-\log u(\theta)] = \theta^{\alpha} p (-\log \theta).$$

By the periodicity of p and \tilde{p} , u satisfies

$$u(\theta e^s) = u(\theta) e^s$$
.

Therefore, if we define $\varphi(\theta) = \tilde{\psi} [u(\theta)]$, it follows that

$$\varphi(\theta W_i) = \tilde{\psi} [u(\theta W_i)] = \tilde{\psi} [u(\theta) W_i],$$

and hence that

$$E \prod_{i=1}^{N} \varphi(\theta W_i) = E \prod_{i=1}^{N} \tilde{\psi} [u(\theta) W_i]$$
$$= \tilde{\psi} [u(\theta)] = \varphi(\theta).$$

Furthermore, since $u(\theta)/\theta$ is bounded away from 0 and ∞ on $(0, \infty)$,

 $\lim_{\theta \downarrow 0} \frac{\log u(\theta)}{\log \theta} = 1.$

Hence

$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta)} = 1 \quad \text{if } v'(\alpha) < 0$$

and

$$\lim_{\theta \downarrow 0} \frac{1 - \varphi(\theta)}{\theta^{\alpha} p(-\log \theta) |\log \theta|} = 1 \quad \text{if } v'(\alpha) = 0,$$

by the corresponding properties of $\tilde{\psi}$. It remains to show that $\varphi \in \mathfrak{L}$. In order to do this, it suffices to note that $T^n \psi \in \mathfrak{L}$ for all n and to use the argument of the proof of Theorem 4.2 to show that

$$\lim_{n\to\infty} T^n \psi = \varphi.$$

The point of the next result is to show that $\mathfrak{P}_{\alpha,s}$ is relatively large for $\alpha \in (0,1)$ and s>0. For simplicity, we will take $s=2\pi$.

(5.2) **Theorem.** Take $\alpha \in (0, 1)$, let a_n and b_n be numbers which satisfy

$$\sum_{n=1}^{\infty} \sqrt{(a_n^2 + b_n^2) \prod_{j=0}^{\infty} \left[1 + \frac{n^2}{(j-\alpha)^2} \right]} \le 1$$

and let

$$p(x) = 1 + \sum_{n=1}^{\infty} \left[a_n \sin nx + b_n \cos nx \right].$$

Then $p \in \mathfrak{P}_{\alpha, 2\pi}$.

Proof. Let D denote differentiation. Then it is easy to check by induction that for $k \ge 1$,

$$D^{k}[\theta^{\alpha}p(-\log\theta)] = \theta^{\alpha-k}(-1)^{k}\left[\prod_{j=0}^{k-1}(D+j-\alpha)p\right](-\log\theta).$$

So, it suffices to show that

$$\left[\prod_{j=0}^{k-1} (D+j-\alpha)p\right](x) \le 0$$

for $k \ge 1$ and $x \in R$. Define $a_n^{(k)}$, $b_n^{(k)}$ and $c^{(k)}$ by

$$\left[\prod_{j=0}^{k-1} (D+j-\alpha)\right] p(x) = -c^{(k)} \left[1 + \sum_{n=1}^{\infty} \left[a_n^{(k)} \sin nx + b_n^{(k)} \cos nx\right]\right].$$

Then $c^{(0)} = -1$, $a_n^{(0)} = a_n$, $b_n^{(0)} = b_n$, $c^{(k+1)} = (k-\alpha)c^{(k)}$,

$$a_n^{(k+1)} = a_n^{(k)} - \frac{n}{k-\alpha} b_n^{(k)}$$

and

$$b_n^{(k+1)} = b_n^{(k)} + \frac{n}{k-\alpha} a_n^{(k)}$$

Therefore $c^{(k)} \ge 0$ for $k \ge 1$ and

$$[a_n^{(k+1)}]^2 + [b_n^{(k+1)}]^2 = \left(1 + \frac{n^2}{(k-\alpha)^2}\right) \{[a_n^{(k)}]^2 + [b_n^{(k)}]^2\},$$

so iterating,

$$[a_n^{(k)}]^2 + [b_n^{(k)}]^2 \le (a_n^2 + b_n^2) \prod_{j=0}^{\infty} \left[1 + \frac{n^2}{(j-\alpha)^2}\right].$$

By the assumption, then,

$$\sum_{n=1}^{\infty} \{ [a_n^{(k)}]^2 + [b_n^{(k)}]^2 \}^{1/2} \le 1$$

for $k \ge 1$. The result then follows from the fact that

$$\sup_{x} |a\sin x + b\cos x| = \sqrt{a^2 + b^2}.$$

The proof of the following theorem follows that of the corresponding result in Kahane and Peyriere (1976).

(5.3) **Theorem.** Suppose v(1) = 0 and v'(1) < 0. Then for any $\mu \in \mathfrak{F}$ and $\beta > 1$,

$$\int_{0}^{\infty} x^{\beta} d\mu < \infty \quad \text{if and only if } v(\beta) < 0.$$

Proof. Let $X, X_1, ..., X_N$ be independent random variables with distribution μ . Suppose first that $EX^{\beta} < \infty$. Then write

$$X^{\beta} \stackrel{\mathfrak{D}}{=} \left(\sum_{i=1}^{N} W_{i} X_{i} \right)^{\beta}$$
$$\geq \sum_{i=1}^{N} W_{i}^{\beta} X_{i}^{\beta},$$

where the inequality is strict on a set of positive probability. Then

$$EX^{\beta} > EX^{\beta} \sum_{i=1}^{N} EW_i^{\beta},$$

so that $v(\beta) < 0$. For the converse suppose $v(\beta) < 0$ and let k be the integer which is determined by $k < \beta \le k+1$. Then for $x_i \ge 0$,

$$\begin{split} \left(\sum_{i=1}^{N} x_{i}\right)^{\beta} & \leq \left(\sum_{i=1}^{N} x_{i}^{\beta/(k+1)}\right)^{k+1} \\ & = \sum_{i=1}^{N} x_{i}^{\beta} + \sum_{i=1}^{N} c_{j_{1},...,j_{N}} (x_{1}^{j_{1}} ... x_{N}^{j_{N}})^{\beta/(k+1)} \end{split}$$

for appropriate constants c_{j_1,\ldots,j_N} . Note that only terms with $\max_i j_i \leq k$ and $\sum_{i=1}^N j_i \leq k+1$ appear in the second summation. Therefore, if Y, Y_1,\ldots,Y_N are independent and identically distributed nonnegative random variables, it follows that

(5.4)
$$E\left(\sum_{i=1}^{N} Y_{i} W_{i}\right)^{\beta} \leq (E Y^{\beta}) e^{v(\beta)} + c \left[E Y^{k}\right]^{\beta/k}$$

for some c which may depend on β but not on the distribution of Y. This can be restated as

$$\int_{0}^{\infty} x^{\beta} d(Tv) \leq e^{v(\beta)} \int_{0}^{\infty} x^{\beta} dv + c \left[\int_{0}^{\infty} x^{k} dv \right]^{\beta/k}$$

for any $v \in \mathfrak{M}$. Recalling from the proof of Theorem 2.7 that

$$\mu = \lim_{n \to \infty} T^n \delta_1$$

where δ_1 is the pointmass at 1, we see that $v(\beta) < 0$ and $\int\limits_0^\infty x^k d\mu < \infty$ imply that $\int\limits_0^\infty x^\beta d\mu < \infty$. Since $v(\beta) < 0$ and v(1) = 0 imply that v < 0 on $(1, \beta]$, the previous statement can be iterated to complete the proof of the theorem. In the first step of the iteration, of course, $\int\limits_0^\infty x d\mu < \infty$ is used, but that is part of the statement of Theorem 2.7.

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Added in Proof. After completion of this paper, B. Mandelbrot suggested that some of the remarks and conjectures in his papers in C.R. Acad. Sci. Paris, Series A (volume 278, pp. 289–292 and 355–358) might be directly relevant to our work. The closest point of contact appears to be his conjecture in paragraph 17, which our results show is incorrect.